

Topics in TCS

Frequency estimation via sketching

Raphaël Clifford

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```
stream  $\langle a_1, \dots, a_m \rangle, a_i \in [n]$   
initialise  $C[1 \dots t][1 \dots k] = 0$   
choose hash functions  $h_1, \dots, h_t : [n] \rightarrow [k]$   
choose hash function  $g_1, \dots, g_t : [n] \rightarrow \{-1, 1\}$ 
```

```
COUNTSKETCH( $a_i$ )  
for each  $j \in [t]$   
     $C[j, h_j(a_i)] += c_j g_j(a_i)$   
  
return  $\hat{f}_{a_i} = \text{median}\{g_j(a_i)C[j, h_j(a_i)]\}$ 
```

c_j is the number of instances of a_i . In the turnstile model this can be either positive or negative.

COUNTSKETCH - worked example



	1	2	3
h_1			
h_2			

	h_1, g_1	h_2, g_2
●	2, +	1, +
●	3, -	2, +
●	1, +	3, -
●	2, -	3, +

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	1	2	3
h_1		+	-
h_2	+	+	

	h_1, g_1	h_2, g_2
●	2, +	1, +
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for each $j \in [t]$
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COUNTSKETCH - worked example



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h_2	++	+	+

	h_1, g_1	h_2, g_2
● (green)	2, +	1, +
● (cyan)	3, -	2, +
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	h_1, g_1	h_2, g_2
	2, +	1, +
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To start, let us look just at an arbitrary row of C . We will show that for each row COUNTSKETCH gives an unbiased estimate. Define $C[x] = C[1, x]$.

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As g and h are independent and g is from a pairwise independent family,

$$\mathbb{E}[g(a)g(j)Y_j] = \mathbb{E}(g(a)) \cdot \mathbb{E}(g(j)) \cdot \mathbb{E}(Y_j) = 0 \cdot 0 \cdot \mathbb{E}(Y_j) = 0$$

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By linearity of expectation

$$\mathbb{E}(X) = f_a + \sum_{j \in [n] \setminus \{a\}} f_j \mathbb{E}[g(a)g(j)Y_j] = f_a$$

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We will now derive the variance of our estimator $X = \hat{f}$. Recall $Y_j = 1$ iff $h(j) = h(a)$.

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$$\begin{aligned}\text{var}(X) &= 0 + \text{var} \left[g(a) \sum_{j \in [n] \setminus \{a\}} f_j \cdot g(j) Y_j \right] \\ &= \mathbb{E} \left[g(a)^2 \sum_{j \in [n] \setminus \{a\}} f_j^2 Y_j^2 + \sum_{\substack{j \in [n] \setminus \{a\} \\ i \neq j}} f_i f_j g(i) g(j) Y_i Y_j \right] - \\ &\quad \left[\sum_{j \in [n] \setminus \{a\}} f_j \mathbb{E}[g(a) g(j) Y_j] \right]^2\end{aligned}$$

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We will need two facts to simplify these terms.

COUNTSKETCH - Analysis IIb

$$\text{var}(X) = \mathbb{E} \left[g(a)^2 \sum_{j \in [n] \setminus \{a\}} f_j^2 Y_j^2 + \sum_{\substack{j \in [n] \setminus \{a\} \\ i \neq j}} f_i f_j g(i) g(j) Y_i Y_j \right] - \left[\sum_{j \in [n] \setminus \{a\}} f_j \mathbb{E}[g(a) g(j) Y_j] \right]^2$$

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Now, the two facts:

1. $\mathbb{E}(Y_j^2) = \mathbb{E}(Y_j) = \Pr(h(j) = h(a)) = \frac{1}{k}$.
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Therefore,

$$\begin{aligned} \text{var}(X) &= \sum_{j \in [n] \setminus \{a\}} \frac{f_j^2}{k} + 0 - 0 \\ &= \frac{\|\mathbf{f}\|_2^2 - f_a^2}{k} \quad \text{where } \mathbf{f} \text{ is the array of frequencies} \end{aligned}$$

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Using the variance $\text{var}(X) = \frac{\|f\|_2^2 - f_a^2}{k}$ we can apply Chebyshev.

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$$\begin{aligned}\Pr(|\hat{f}_a - f_a| \geq \epsilon \sqrt{\|\mathbf{f}\|_2^2 - f_a^2}) &= \Pr(|X - \mathbb{E}(X)| \geq \epsilon \sqrt{\|\mathbf{f}\|_2^2 - f_a^2}) \\ &\leq \frac{\text{var}(X)}{\epsilon^2(\|\mathbf{f}\|_2^2 - f_a^2)} \\ &= \frac{1}{k\epsilon^2} \\ &= \frac{1}{3} \quad (\text{set } k = 3/\epsilon^2)\end{aligned}$$

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$$\begin{aligned}\Pr(|\hat{f}_a - f_a| \geq \epsilon \sqrt{\|\mathbf{f}\|_2^2 - f_a^2}) &= \Pr(|X - \mathbb{E}(X)| \geq \epsilon \sqrt{\|\mathbf{f}\|_2^2 - f_a^2}) \\ &\leq \frac{\text{var}(X)}{\epsilon^2(\|\mathbf{f}\|_2^2 - f_a^2)} \\ &= \frac{1}{k\epsilon^2} \\ &= \frac{1}{3} \quad (\text{set } k = 3/\epsilon^2)\end{aligned}$$

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$$\Pr(|\hat{f}_a - f_a| \geq \epsilon \|\mathbf{f}_{-a}\|_2) \leq \frac{1}{3}$$

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For an arbitrary token a , the probability of being further than $\epsilon \|\mathbf{f}_{-a}\|_2$ from the correct frequency is at most $\exp(-t/36)$.

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Running time: one-pass and $O(t)$ time per token.

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Assuming we set $k = 3/\epsilon^2$, for an arbitrary token a , the probability that COUNTSKETCH's estimate is further than $\epsilon \|\mathbf{f}_{-a}\|_2$ from the correct frequency is at most $\exp(-t/36)$.

COUNT-MIN sketch

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```
stream  $\langle a_1, \dots, a_m \rangle, a_i \in [n]$   
initialise  $C[1 \dots t][1 \dots k] = 0$   
choose hash functions  $h_1, \dots, h_t : [n] \rightarrow [k]$ 
```

```
COUNT-MIN( $a_i$ )  
for each  $j \in [t]$   
     $C[j, h_j(a_i)] += c_i$ 
```

```
return  $\hat{f}_a = \min_{1 \leq i \leq t} C[i, h_i(a)]$ 
```

c_i is the number of instances of a_i . In the turnstile model this can be either positive or negative.

COUNT-MIN - worked example

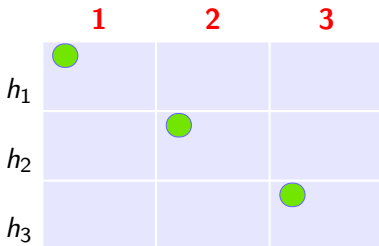


	1	2	3
h_1			
h_2			
h_3			

	h_1	h_2	h_3
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●	2	1	1
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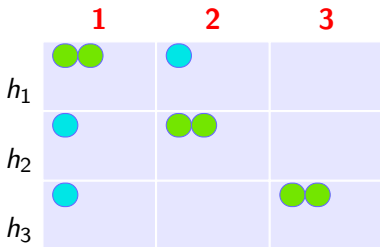


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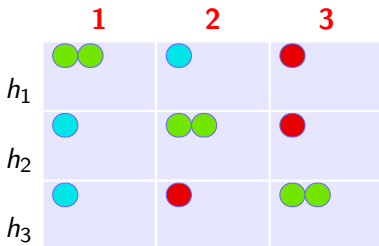
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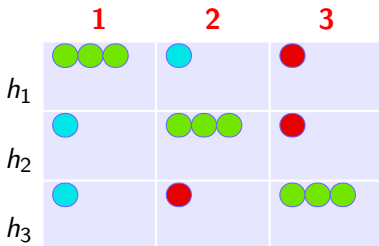
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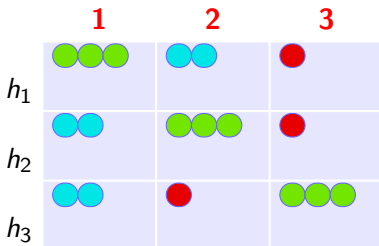
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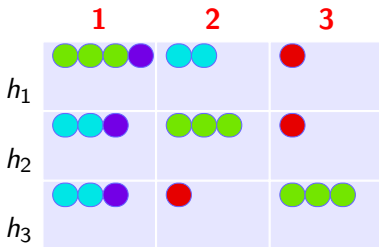
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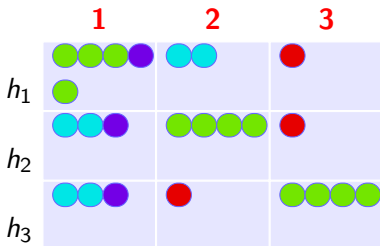

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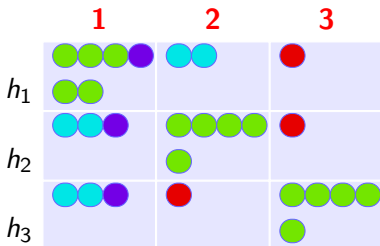
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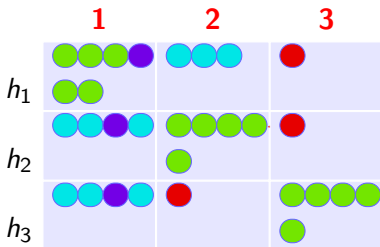
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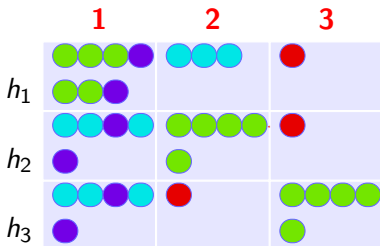
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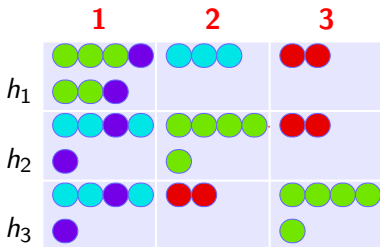


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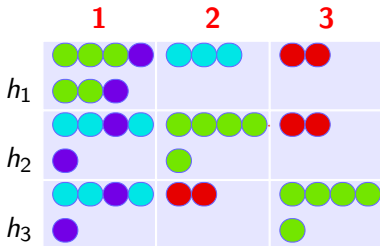
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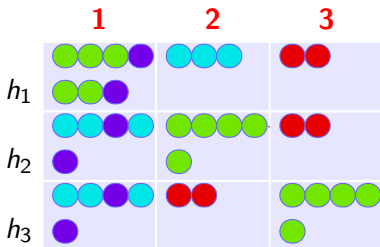


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COUNT-MIN - worked example



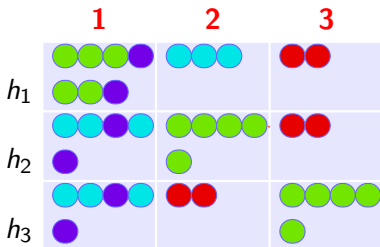
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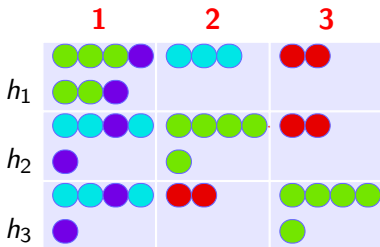
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By Markov's inequality

$$\Pr(X_i \geq \epsilon \|\mathbf{f}_{-a}\|_1) \leq \frac{\|\mathbf{f}_{-a}\|_1}{k\epsilon \|\mathbf{f}_{-a}\|_1} = \frac{1}{2} \quad \text{set } k = 2/\epsilon$$

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We have a bound for a single counter. Over t counters the reported excess is the minimum over all X_j . We can now derive the probability that all the excesses are at least $\epsilon \|\mathbf{f}_{-a}\|_1$ directly.

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The space usage is better than COUNTSKETCH by a factor of $1/\epsilon$. COUNT-MIN's error probability is bounded by $\epsilon \|\mathbf{f}_{-a}\|_1$ instead of $\epsilon \|\mathbf{f}_{-a}\|_2$ for COUNTSKETCH.

COUNT-MIN - Analysis II

We have a bound for a single counter. Over t counters the reported excess is the minimum over all X_j . We can now derive the probability that all the excesses are at least $\epsilon \|\mathbf{f}_{-a}\|_1$ directly.

$$\Pr(\hat{f}_a - f_a \geq \epsilon \|\mathbf{f}_{-a}\|_1) \leq \frac{1}{2^t} = \delta \quad \text{set } t = \left\lceil \log_2\left(\frac{1}{\delta}\right) \right\rceil$$

$k = 2/\epsilon, t = \lceil \log_2(1/\delta) \rceil$ gives total space in bits

$$O\left(\frac{1}{\epsilon} \log \frac{1}{\delta} \cdot (\log m + \log n)\right)$$

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COUNTSKETCH: with $k = \lceil 3/\epsilon^2 \rceil$ and $t = \lceil \log_2(1/\delta) \rceil$,

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MISRA-GRIES uses $O((1/\epsilon)(\log m + \log n))$ bits but does not work in the turnstile model (with deletions).